

## STOCHASTIC PROCESSES 29-01-2021

2 hours 30 minutes. No documents allowed.

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### 1. EXERCISE 1

The two parts of this first exercise are (essentially) independent.

**Part 1.**  $X$  is a  $Q$ -MC on  $E = \mathbb{N} = \{1, 2, \dots\}$  with  $Q$  defined by

$$Q(k, k+1) = \frac{k}{k+1} \quad \text{and} \quad Q(k, 1) = \frac{1}{k+1}.$$

- (1) Explain, in a concise but complete way, why this chain is irreducible and aperiodic.
- (2) Show that this chain is null recurrent.
- (3) Compute  $\lim_{n \rightarrow \infty} \mathbb{P}(X_n = 2) / \mathbb{P}(X_n = 3)$ .

**Part 2.**  $X$  is now a  $Q$ -MC on  $E = \mathbb{N} = \{1, 2, \dots\}$  with  $Q$  defined by

$$Q(k, k+1) = \frac{k}{k+2} \quad \text{and} \quad Q(k, 1) = \frac{2}{k+2}.$$

Like in the first part, the chain is irreducible and aperiodic.

- (4) Show that this chain is positive recurrent. We recall that  $\sum_{k=2}^{\infty} 1/(k(k+1)(k+2)) = 1/12$  (which may or may not be needed).
- (5) We call  $\pi$  the invariant probability for the chain: show that  $\mathbb{E}_{\pi}[X_n] = \infty$  for all  $n$ .
- (6) With  $T_k = \inf\{n = 1, 2, \dots : X_n = k\}$ , compute  $\mathbb{P}_1(T_1 = n)$  and (verify with)  $\mathbb{E}_1[T_1]$ .
- (7) ( $\star$ ) We show now that,  $\mathbb{P}_{\pi}$ -a.s.,  $\limsup_n X_n/n = 0$ .<sup>1</sup>
  - (a) Set  $\tau_j$  the  $j^{\text{th}}$  visit to 1 by the chain and  $\tau_0 := 0$ . Explain briefly why  $\tau_j$  is a stopping time and (still briefly) why  $(\tau_j - \tau_{j-1})_{j=1,2,\dots}$  is an independent sequence with  $(\tau_j - \tau_{j-1})_{j=2,3,\dots}$  is IID. Explain why the Law of Large Numbers implies  $\lim_j (\tau_j - \tau_{j-1})/j = 0$  a.s..
  - (b) For  $n$  set  $j_n = j_n(\omega)$  such that  $\tau_{j_n-1} < n \leq \tau_{j_n}$ . Prove that a.s.  $j_n \sim n/\mathbb{E}_1[\tau_1]$  and then conclude.

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### 2. EXERCISE 2

In the first part we consider the classical urn of Polya: we start with an urn that contains  $R_0 > 0$  red balls and  $B_0 > 0$  blue balls. At each time step we choose one ball from the urn and we put it back together with  $\alpha > 0$  balls of the same color. Let  $R_n$  and  $B_n$  the number of red and blue balls at step  $n$ . So,  $(R_{n+1}, B_{n+1}) = (R_n + \alpha, B_n)$  with probability  $R_n/(R_n + B_n)$  and  $(R_{n+1}, B_{n+1}) = (R_n, B_n + \alpha)$  with probability  $B_n/(R_n + B_n)$ . This defines the the two sequences  $(R_n)$  and  $(B_n)$  of random variables: of course  $R_n + B_n = R_0 + B_0 + \alpha n$ . We set  $\mathcal{F}_n := \sigma(R_1, \dots, R_n)$ . We consider non random  $(R_0, B_0)$ , so we choose  $\mathcal{F}_0$  trivial.

- (1) Give the definition of Uniform Integrability (UI) for a sequence of random variables. Explain, by stating the appropriate theorems, why if a UI sequence  $(X_n)$  converges in law to  $X$ , then  $\lim_n \mathbb{E}[X_n] = \mathbb{E}[X]$ .

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<sup>1</sup>What is obvious is that  $\limsup_n X_n/n \leq 1$  (because  $X_n \leq X_0 + n$ ). This fact is of interest because it shows that a stationary MC can be very different from an IID sequence: in the IID case we know that  $\mathbb{E}[|X_1|] = \infty$  if and only if  $\limsup_n X_n/n = 0$ .

- (2) Show that  $(M_n)_{n=0,1,\dots}$ , with  $M_n = R_n/(R_n + B_n)$  is an UI martingale.  
 (3) Explain why  $M = \lim_n M_n$  exists a.s. and in  $\mathbb{L}^1$ . Explain also (by at least sketching the proof given in the course) why  $M_n = \mathbb{E}[M|\mathcal{F}_n]$ . Deduce that  $M$  is not a constant.  
 (4) Show that  $(Y_n)_{n=0,1,\dots}$ , with  $Y_n = R_n(R_n + \alpha)/((R_n + B_n)(R_n + B_n + \alpha))$ , is also an UI martingale  
 (5) Compute the variance of  $M$  and conclude, once again, that  $M$  is not a constant.  
*Obs.: in case you want to check your result, the law of  $M$  is known and for  $R_0 = B_0 = \alpha = 1$  it is  $U(0, 1)$ , whose variance is  $1/12$ .*

Now we *slightly* generalize the model (*urn of Friedman*): when putting back the ball we have chosen together with  $\alpha$  balls of the same color, we put also  $\beta$  balls of the other color (the general case can be treated, but choose  $\beta < \alpha$ ). We define this way two new sequences  $(R_n)$  and  $(B_n)$  of random variables: this time  $R_n + B_n = R_0 + B_0 + (\alpha + \beta)n$ . The filtration we choose is once again defined by  $\mathcal{F}_n := \sigma(R_1, \dots, R_n)$ . We set

$$\delta := \alpha - \beta, \quad \tau := \alpha + \beta \quad \text{and} \quad \rho := \frac{\delta}{\tau}.$$

For what follows you may want to use that for  $a, b, c > 0$  we have

$$\prod_{j=0}^n \left(1 + \frac{a}{b + cj}\right) \stackrel{n \rightarrow \infty}{\sim} C_{a,b,c} n^{a/c}.$$

where<sup>2</sup>  $C_{a,b,c} > 0$ . As usual,  $f(n) \sim g(n)$  means  $\lim_n f(n)/g(n) = 1$ .

- (6) Set  $Z_0 := R_0 - B_0$  and

$$Z_n := (R_n - B_n) / \prod_{j=0}^{n-1} \left(1 + \frac{\delta}{R_0 + B_0 + \tau j}\right),$$

for  $n = 1, 2, \dots$ . Compute  $\mathbb{E}[R_{n+1} - B_{n+1}|\mathcal{F}_n]$  and infer that  $(Z_n)$  is a martingale.

- (7) Show that

$$\mathbb{E}[(R_{n+1} - B_{n+1})^2] = \left(1 + \frac{2\delta}{R_0 + B_0 + n\tau}\right) \mathbb{E}[(R_n - B_n)^2] + \delta^2 =: a_n \mathbb{E}[(R_n - B_n)^2] + \delta^2,$$

that is (no need to show it, no probability involved)

$$\mathbb{E}[(R_n - B_n)^2] = \left((R_0 - B_0)^2 + \delta^2 \sum_{j=0}^{n-1} \frac{1}{\prod_{k=0}^{j-1} a_k}\right) \prod_{j=0}^{n-1} a_j,$$

where empty sums should be read as zero, as well as empty products should be read as one.

- (8) Show that if  $\rho > 1/2$  we have  $\sup_n \mathbb{E}[Z_n^2] < \infty$ .  
 (9) Explain why we therefore have (if  $\rho > 1/2$ ) that  $\lim_n (R_n - B_n)/n^\rho$  exists a.s. and in  $\mathbb{L}^1$ .

*Conclusion and curiosities:* let us admit it, the result is extremely surprising. The Friedman urn seems a minor modification of the Polya urn, but point (9) says that  $\lim_n R_n/(R_n + B_n) = 1/2$  a.s.. In the Polya urn  $\lim_n R_n/(R_n + B_n)$  is a non trivial random variable!

The fact that  $\lim_n R_n/(R_n + B_n) = 1/2$  a.s. is true for every  $\rho \in (0, 1)$  and this general result can be obtained by arguments similar to the one we just used. For more on this: D. A. Freedman, *Bernard Friedman's urn*, Ann. Math. Statist. 36 (1965), 956-970.

<sup>2</sup>Actually,  $C_{a,b,c} = \Gamma(b/c)/\Gamma((a+b)/c)$  but you will not need that.